statistics cheat sheet

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1 my first cheat sheet

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Chapter one, Probability Definitions:

1) 2 events are disjoint if they have no outcomes in common. Written as A∩B=Ø

2) 2 events are independent if occurrence of one does not give any indication of the occurrence of the other, in addition, we write $P(A \cap B)=P(A) P(B)$

3) Sample space Ω={A,B,....} contains all possible events



Permutation and Combination: In Permutation the order we list things is important. Hence A,B would be different than B,A. Hence result of Permutations are *larger* than combination.



Figure 1:²statistics

2 second cheat sheet

problem: phone calls received at rate $\lambda = 2$ per hr. If person wants to take 10 min shower, what is probability a phone will ring during that time?

answer: first change to $\omega = \lambda_{60}^{10} = 2\frac{10}{60} = .3333$, now we want $P(X \ge 1) = 1 - P(X \le 1) = 1 - P(0)$

but $P(k) = \frac{\lambda^k}{k!}e^{-\lambda}$, but remember, we are using ω , so $P(k) = \frac{\omega^k}{k!}e^{-\omega}$ so $P(0) = \frac{.3333^0}{0!}e^{-.3333} = 0.777$

so $P(X \ge 1) = 1 - .777 = 0.283$, so 28% change the phone will ring.

How long can shower be if they wish probability of receiving no phone calls to be at most 0.5?

 $P(0)=0.5=\frac{\omega0}{0!}e^{-\omega}\rightarrow 0.5=e^{-\omega}$ hence $\ln 0.5=-\omega\rightarrow\omega=0.693,$ so $\lambda\frac{x}{60}=0.693\rightarrow x=20.7$ minutes

To find quantile, say $\frac{1}{4}$, first find an expression for F(x) as function of x, then solve for x in F(x) = .25

For median, solve for x in F(x) = .5

properties of CDF: 1. Show $F(x) \ge 0$ for all x. Do this by showing $F'(x) \ge 0$, and show limit $F(x) \to 1$ as $x \to \infty$ and limit $F(x) \to 0$ as $x \to -\infty$. And $P(k_1 \le T < k_2) = F(k_2) - F(k_1)$

properties of pdf:

- 1. piecewise continuous
- 2. $pdf(x) \ge 0$
- 3. $\int_{-\infty}^{\infty} p df(x) = 1$

remember $\frac{d}{dx} \tan^{-1} x = \frac{1}{1+x^2}$

The geometric distribution is the only discrete memoryless random distribution. It is a discrete analog of the exponential distribution. continuous

Some relations

$$\sum_{k=1}^n k = \frac{1}{2}n(n+1)$$

Geometric sum

$$\sum_{k=0}^n r^k = \frac{1-r^{n+1}}{1-r}$$

if -1 < r < 1, then

$${\displaystyle \sum_{k=0}^{\infty}} r^k = \frac{1}{1-r}$$

if the sum is from 1 then

$$\sum_{k=1}^{n} r^{k} = \frac{r(1-r^{n+1})}{1-r}$$

if -1 < r < 1, then

$$\sum_{k=1}^{\infty} r^k = \frac{r}{1-r}$$

$$\Gamma(x) = \int_0^{\infty} u^{x-1} e^{-u} du$$

$$\Gamma(n) = (n-1)!$$

$$\frac{d}{dx} \ln(x) = \frac{1}{x}$$

$$\int \ln(y) \, dy = -y + y \ln(y)$$

$$\int \frac{1}{y} dy = \ln(y)$$

And

$$egin{pmatrix} n \ n_1 & n_2 & n_3 \end{pmatrix} = rac{n!}{n_1! \; n_2! \; n_3!}$$

If given joint density $f_{XY}(x,y)$ and asked to find conditional $P(X|Y) = \frac{f_{XY}(x,y)}{f_Y(y)}$ so need to find marginals. Marginal is found from $f_Y(y) = \int_x f_{XY}(x,y) \, dx$, and $f_X(x) = \int_y f_{XY}(x,y) \, dy$

To convert from x, y to polar, example: given $f(x, y) = c\sqrt{1 - (x^2 + y^2)}$ find c, where $x^2 + y^2 \le 1$, then write

$$c\int_{\theta=-\pi}^{\theta=\pi}\int_{r=0}^{r=1}\sqrt{1-r^2}rdrd\theta$$

Use identity above.

law of total probablity: if we know Y|X and X and want to know distribution of Y, then $f(Y) = \int_{-\infty}^{\infty} f_{Y|X}(y|x) f_X(x) dx$

$$\sum_{n_{1}}^{n} \sum_{n_{2}}^{n} \sum_{n_{3}}^{n} \sum_{n_{3}}^{n} \sum_{n_{3}}^{n} \sum_{n_{3}}^{n} \sum_{n_{1}, n_{2}, n_{3}}^{n} \sum_{n_{1}, n_{2}, n_{3}}^{n}$$

 $y < 0, f_T(y) = \int_{x \to y}^{x \to y} f(x, y) dx$ and leave it at that. do not Add them.

Figure 2: multi

$$Z = \frac{\bar{X}_n - \mu}{\sigma/\sqrt{n}} \to N(0, 1)$$
$$T = \frac{\bar{X}_n - \mu}{S_n/\sqrt{n}} \to T(n)$$

where S_n is *std* of the sample.

Note Var(sample) has chi square (n) distribution. CI for T:

$$\Pr\left(-A < \frac{\bar{X}_n - \mu}{S_n / \sqrt{n}} < A\right) = 1 - \alpha$$
$$\Pr\left(\bar{X}_n - A\frac{S_n}{\sqrt{n}} < \mu < \bar{X}_n + A\frac{S_n}{\sqrt{n}}\right) = 1 - \alpha$$